

**Original article:**

**POSTMENOPAUSAL HORMONE AND THE RISK OF  
NEPHROLITHIASIS: A META-ANALYSIS**

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<http://dx.doi.org/10.17179/excli2017-304>

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**ABSTRACT**

Menopause is reported to be associated with increased urinary calcium excretion, which may enhance the risk for the development of calcium kidney stones. However, it remains controversial about whether high level of postmenopausal hormone (PMH) is a risk factor for formation of nephrolithiasis. Several observational studies have shown that PMH is protective based on 24-hour urinary parameters. Recent clinical trials provided evidence to conclude that estrogen therapy increases the risk of nephrolithiasis in healthy postmenopausal women. Our study aimed to comprehensively assess clinical evidence on the relationship between postmenopausal hormone level and risk of nephrolithiasis. To conduct systematic review, we pooled total 98 potentially related articles in Cochrane library, Medline, and Embase. Three studies with a total of 71101 study participants that included two clinical trials, 4 stratified and potentially usable results by the status of menopause and type of PMH use derived from one prospective cohort study, and one case-control studies were selected to pool relative risk using random-effect model. How the difference in menopause status, whether naturally menopausal or surgically menopausal, influenced the pooled relative risk was included in the subgroup analysis. The study population aged from 45 to 70 years old. The follow-up year and adjusted confounders differed across different studies. The pooled relative risk for the 7 stratified studies was 0.91 (95 % confidence interval (CI): [0.72, 1.14]). In the menopausal status-specific analysis, the pooled relative risk for naturally menopausal women was 0.92 (95 % CI, [0.64, 1.27];  $I^2 = 82.74\%$ ) whereas the pooled relative risk for surgically postmenopausal women is 0.90 (95 % CI, [0.63, 1.29];  $I^2 = 78.47\%$ ). The above results suggested that there was no significant association between PMH and the risk of nephrolithiasis. The difference in menopausal status did not influence the relationship between PMH and the risk of kidney stone formation.

**Keywords:** postmenopause, nephrolithiasis, hormone, women, kidney stone

**INTRODUCTION**

Nephrolithiasis (kidney stone) is a common disease among postmenopausal women, affecting around 5 %-7 % of the population in the United States (Novak et al., 2009). A

sharp increase in urinary calcium stone formation after menopause implies a close correlation between hormone level and pathology of nephrolithiasis. Estrogen replacement, which is aimed at relieving symptoms of men-

opause, increased citrate and calcium excretion in rates in postmenopausal women with recurrent urolithiasis (Dey et al., 2002). The increased stone inhibitory citrate level and increased agglomeration inhibition by estrogen replacement implied an overall beneficial and protective effect of estrogen on the risk of calcium oxalate stone formation in postmenopausal women (Iguchi et al., 1999). However, recent studies found that hormone use such as estrogen replacement therapy among healthy women may be associated with a higher risk of nephrolithiasis (Maalouf et al., 2010; Mattix Kramer et al., 2003). The seemingly-conflicted results may be due to the diverse nature of studies conducted in order to investigate how hormone is associated with the risk of kidney stone formation. Different cohort samples may also contribute to differences in the results. To the best of our knowledge, there is very few systematic review on hormone use and the risk of nephrolithiasis among healthy postmenopausal women. Here we conducted a systematic review and prospective meta-analysis on several observational studies and clinical trials to examine whether the postmenopausal hormone is a risk factor for kidney stone formation. We aimed at exploring the true effect of PMH on the risk of incident kidney stones.

## MATERIAL AND METHODS

### *Searching strategy and selection criteria*

Literature searching was performed by using Cochrane library, Medline, and Embase. The searching terms included combinations of keywords and their synonyms, such as “Nephrolithiasis”, “postmenopausal”, “Estrogen”, “kidney stone”, “renal lithiasis”, “renal stone”, “women”, “observational”, and “Clinical Trials”. And the search strategies used for the other databases were similar, with the necessary adaptations made. Additionally, we manually searched reference lists in the selected studies to identify potentially relevant studies.

### *Study selection and data extraction*

We set up the inclusion and exclusion criteria to include eligible studies. We focused on postmenopausal women aged around 45 to 79 years with no previous history of kidney stones and selected cohort studies that compared the risk of kidney stones among the group with PMH treatment and the group without. The estimated values of odds ratio (OR), relative risk (RR) or hazard ratio (HR) and its 95 % confidence interval (CI) needed to be specifically reported. Demographic characteristics such as age, sex, body mass index (BMI), serum calcium, and urinary calcium were described. Confounders such as age were adjusted if necessary. The exclusion criteria included: (1) only age-adjusted or other confounder-adjusted OR, RR, or HR were reported. (2) The subjects of the study had recurrent nephrolithiasis or other significant chronic kidney diseases. (3) The data were extracted from the same study population. After carefully reviewing the selected studies, the following information on the publications was collected: abstract, full text, title, author information (i.e. first name, last name, initials), and publication year. Characteristics included country of origin, sample size, age, follow-up time, type of risks and confounders-adjusted risks were extracted from the selected studies.

### *Statistical analyses*

Multivariate-adjusted outcome data (ORs, RRs, HRs and 95 % CI) were extracted and converted logarithmically to log (RR) for each study. The I-squared statistic was used for heterogeneity test, and the log (RR) of the studies were pooled using fixed effect models when the heterogeneity is low (i.e., I-squared < 30%). Otherwise, a random effect model was used. Forest plots were constructed to visually demonstrate RRs and their confidence interval. Funnel plot was constructed after Trim-and-Hill correction for evaluating funnel asymmetry. Fail-safe N was used to determine the number of NULL studies that have to be added to reduce the significance of the meta-analysis to (0.05). Publication bias

was analyzed by using Egger's regression test. The sensitivity analyses were conducted by omitting one study at a time to recalculate the pooled RR. All above analyses were performed by R version 3.3.2 with functional packages including ggplot2, metafor, rmeta, and meta.

## RESULTS

### *Literature search*

A total of 98 potentially related articles were identified after employing the searching strategy in Medline (PubMed), Embase, and Cochrane library. 76 duplicated articles or non-research type's articles or non-English written articles were removed after screening their titles and abstracts. The remaining 22 articles were carefully reviewed by several independent reviewers. The selection process was demonstrated as shown in Figure 1.

### *Study characteristics*

Three studies including one clinical study, one prospective cohort study, and case-control study were included in the meta-analysis. The characteristics of publications and their included demographics were summarized as

shown in Table 1. Two studies were conducted in the United States and one case-control study was conducted in China. The study population was aimed at postmenopausal women with natural or surgical menopause and matched-up controls. Mattix Kramer et al. (2003) studied on postmenopausal women with 26,251 undergoing natural menopause and 17,306 undergoing surgical menopause: mean age and BMI of natural postmenopausal females were 60.4 years old and 25.7 kg/m<sup>2</sup>, while mean age and BMI of surgical postmenopausal females were 58.3 and 25.8. The study by Maalouf et al. (2010) contains two clinical trials: 10,739 postmenopausal women with hysterectomy were enrolled in the estrogen-alone trial, while 16,608 postmenopausal women without hysterectomy were enrolled in the estrogen plus progestin (E+P) trial. Women in the estrogen-alone trial were randomized to receive 0.625 mg/d of conjugated equine estrogens (CEE) or matching placebo. Women in the E+P trial were given a single tablet of CEE plus 2.5 mg/d of medroxyprogesterone acetate or matching placebo. Mean age and BMI for participants without hysterectomy in E+P trial were around

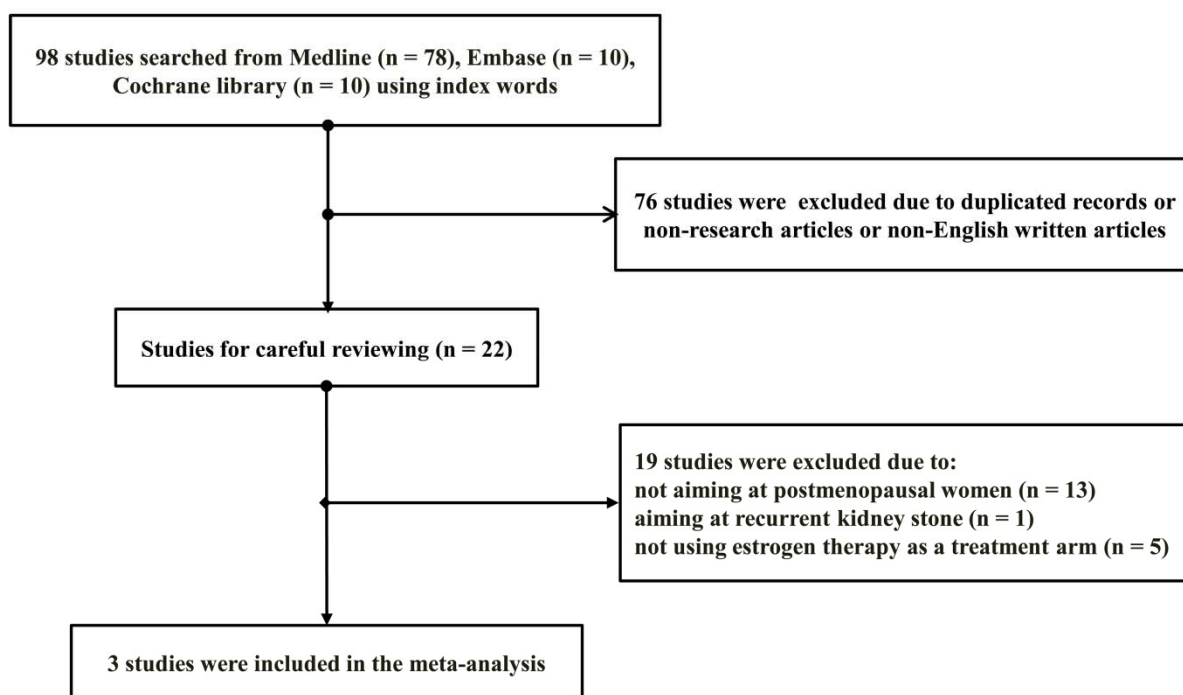


Figure 1: Process of study selection

**Table 1:** Characteristics of the selected studies

	<b>Mattix Kramer et al., 2003</b>	<b>Maalouf et al., 2010</b>	<b>Zhao et al., 2013</b>
<b>Study design</b>	Prospective cohort study	Randomized clinical trials	Case-control study
<b>Country</b>	United States	United States	China
<b>Study Population</b>	43,557	27,347	197
<b>Natural Post-menopausal</b>	26,251	16,608	113
<b>Surgical Post-menopausal</b>	17,306	10,739	/
<b>Age/years * (Natural)</b>	60.4 (5.4)	63.6	48-69
<b>BMI/ kg/m<sup>2</sup>* (Natural)</b>	25.7 (4.9)	30.1	15.8-32.5
<b>Age/years * (Surgical)</b>	58.3 (6.6)	63.3	/
<b>BMI/ kg/m<sup>2</sup>* (Surgical)</b>	25.8 (4.9)	28.5	/
<b>Adjusted confounders</b>	age, body mass index, presence of hypertension, age at menopause, supplemental calcium, alcohol consumption, diet, and fluid intake	Age, race, diet, BMI, and medical history	demographic data and medical history
<b>follow-up/years</b>	18	5.6	/
<b>Type of risk</b>	RR	HR	OR

\* Age and BMI were expressed differently across the studies: mean, mean (SD), or range was reported individually.

60.3 years old and 28.5 kg/m<sup>2</sup> while mean age and BMI for participants with hysterectomy in the estrogen-alone trial were around 63.6 years old and 30.1 kg/m<sup>2</sup>. All participants, including total 113 female patients with newly diagnosed kidney stones after menopause and 84 age frequency-matched stone-free female controls, enrolled in the case-control study conducted by Zhao et al. (2013) were naturally postmenopausal: The range of age of the study subjects was 48-69 while the range of BMI of the study subjects was 15.8-32.5; Odds ratios (ORs) for associations between sex hormones indicated by the level of serum testosterone (T) and estradiol (E2) and kidney stones were estimated with logistic regression models. Potential confounding adjusted factors slightly differed across the three studies, and the primarily adjusted factors

were demographic data such as age, BMI, diet, and medical history.

### Primary analysis

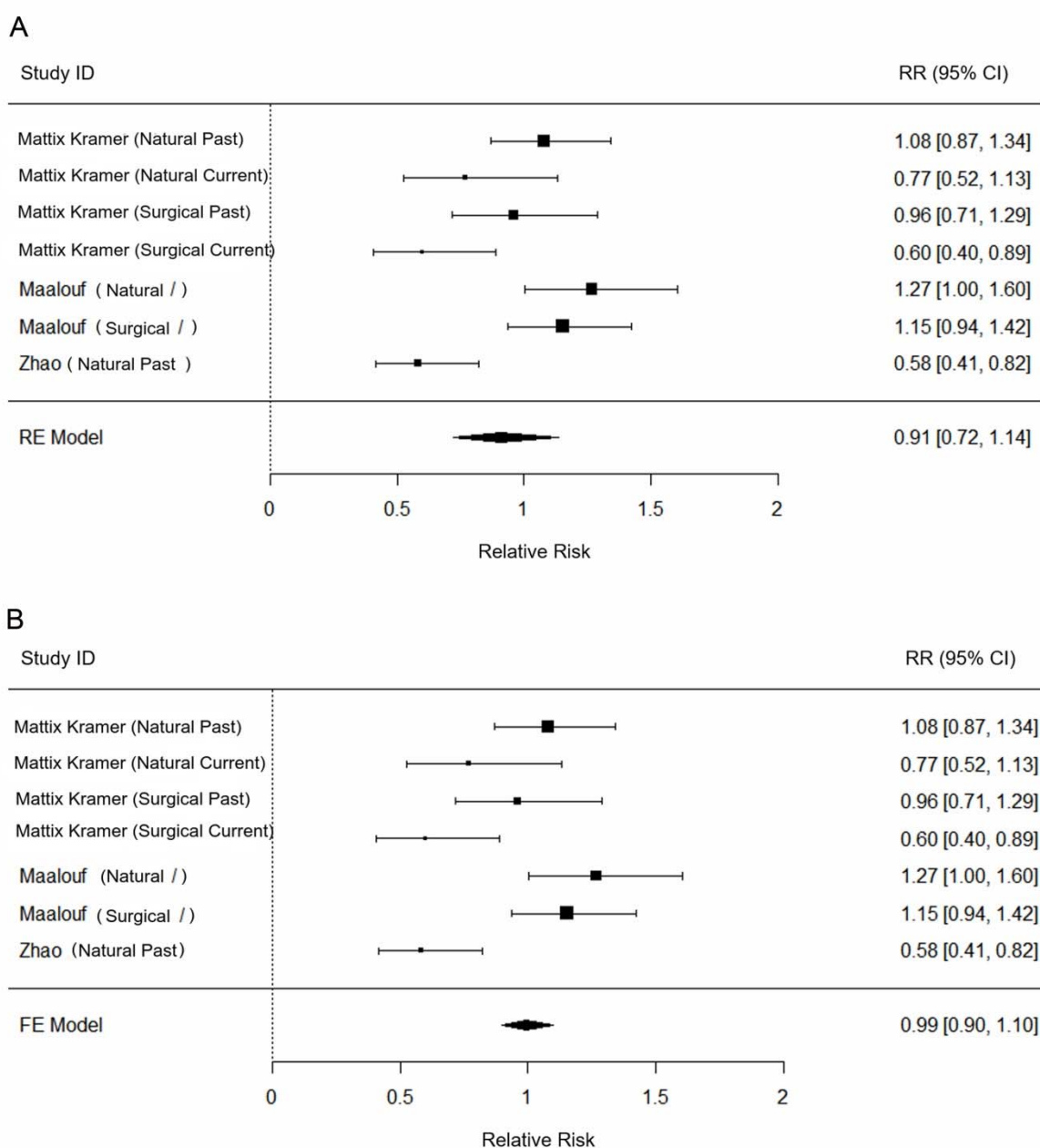
To investigate the association between PMH and the incident of kidney stone, we first stratified the results of study population in Mattix Kramer et al. (2003), into two kinds which target naturally postmenopausal and surgical postmenopausal women, in which each kind of study subjects were further divided into past users and current users with a duration ranged from 5 to 9.9 years. For the two clinical trials in Maalouf et al. (2010), the results were separately analyzed. Fixed effects or random effects modeling were performed on the extracted relative risks (Figure 2A-B). Through heterogeneity test, we found

$I^2$  is 78.22 %, suggesting a high inter- and intra-study variation. Thus, a random effect modeling was used to perform the meta-analysis. According to the result analyzed by random effect modeling (**Figure 2A**), there is no significant association between PMH and the risk of kidney stone formation (the pooled relative risk: 0.91, confidence interval (CI): [0.72, 1.14],  $p > 0.05$ ).

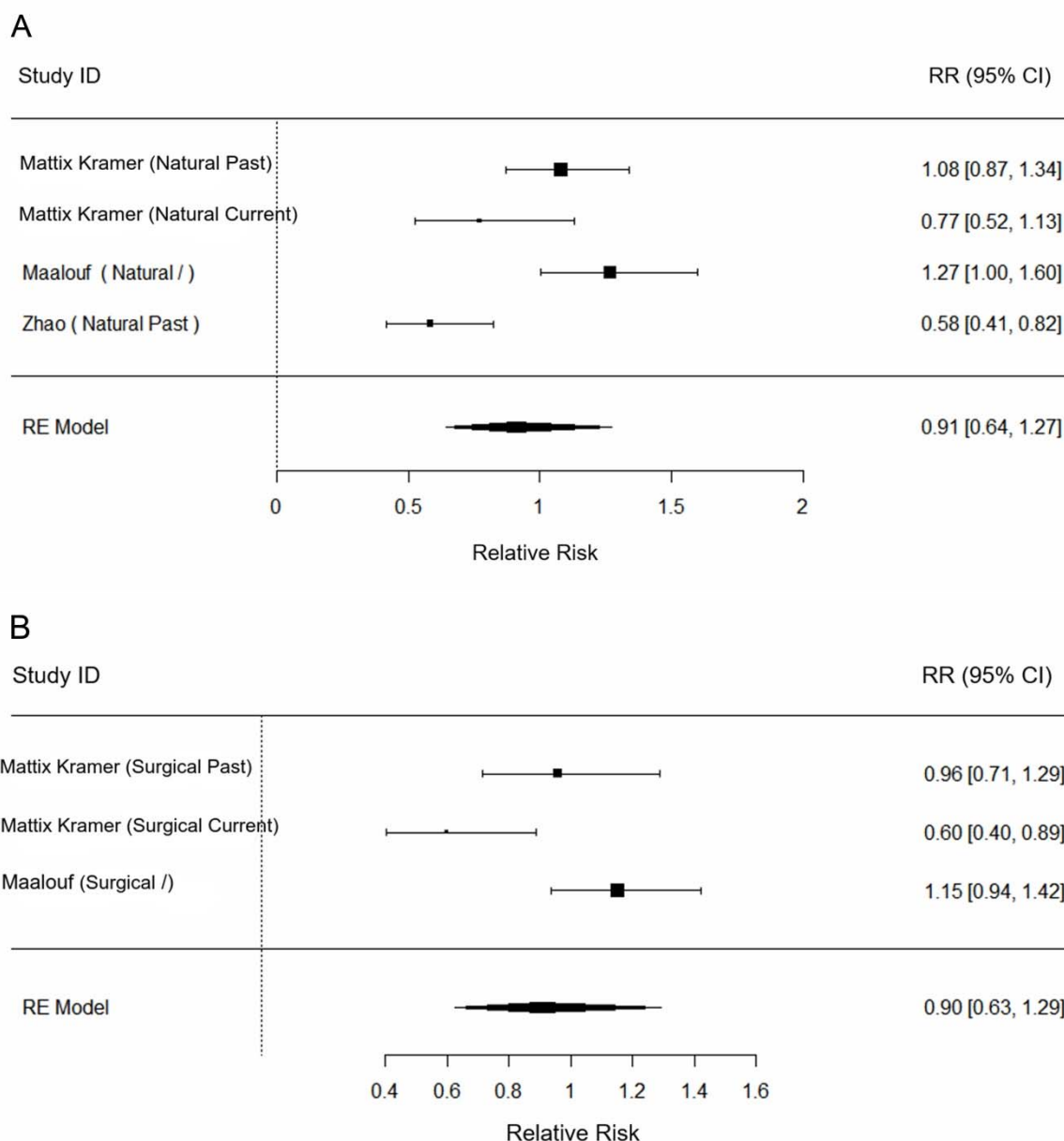
### Subgroup analysis

In our subgroup analysis, we aimed at investigating whether different menopause

types influenced the association between PMH and the risk of nephrolithiasis. According to the results (**Figure 3A-B**), neither significant association between PMH and risk of kidney stone formation was found in women with natural menopause or with surgical menopause (Naturally postmenopausal women: pooled RR, 0.91; 95 % CI, [0.64, 1.27];  $I^2 = 82.74$  %; Surgically postmenopausal women: pooled RR, 0.90; 95 % CI, [0.63, 1.29];  $I^2 = 78.47$  %).



**Figure 2:** Forest plot of the association between postmenopausal hormone and incident kidney stone by random effect (RE) model (**A**) and fixed effect (FE) model (**B**).



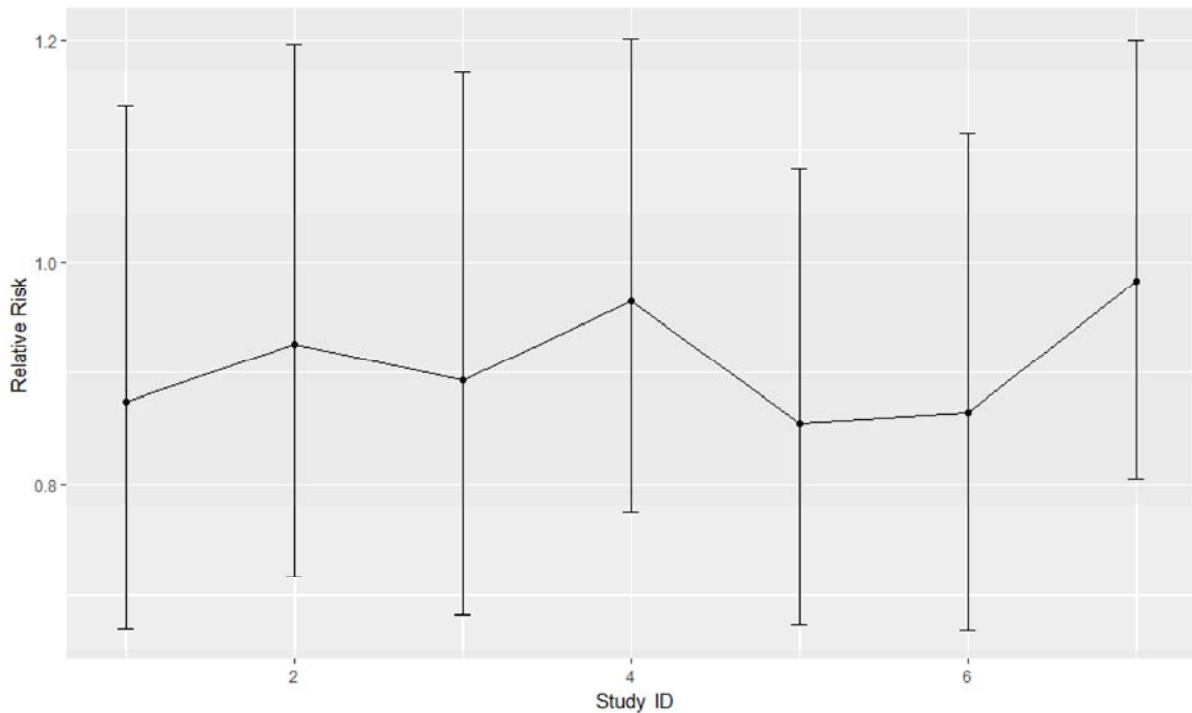
**Figure 3:** Forest plot of the association between postmenopausal hormone (PMH) and incident kidney stone by different types of menopause (**A, B**).

### Sensitivity analysis

To analyze sensitivity for each different stratified study, we conducted sensitivity analysis by omitting one stratified study and re-calculating pooled relative risk (Figure 4). Based on the result, the primary results were not influenced by omitting one study at a time.

### DISCUSSION

In this meta-analysis including 7 stratified results and a total of 71,101 participants, we demonstrated that PMH is not associated with a statistically significant increased or decreased risk of nephrolithiasis. Different types of menopause (surgical or natural) did not contribute to identifying significantly increased or decreased risk of kidney stone.

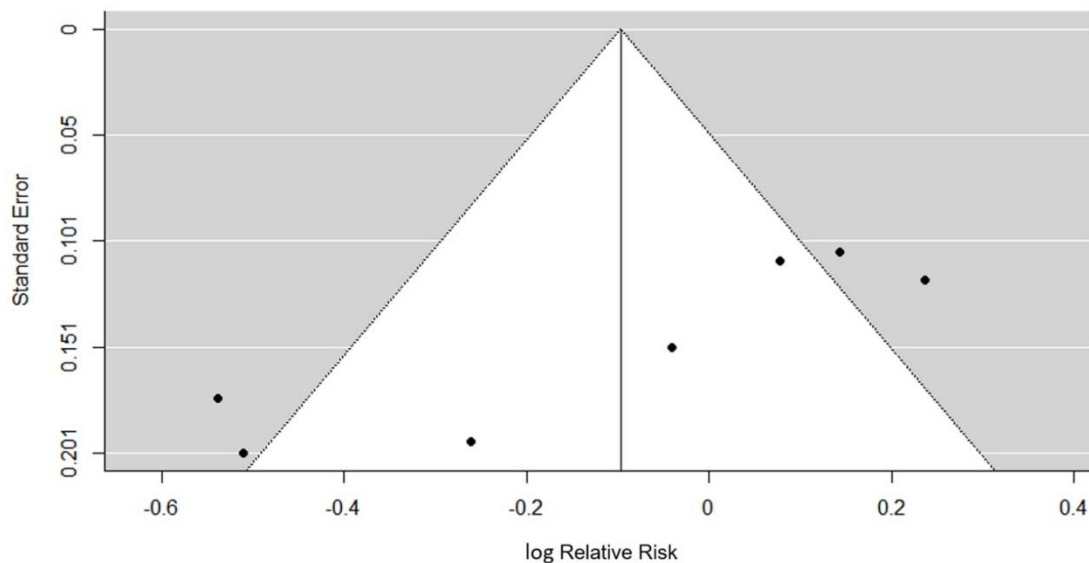


**Figure 4:** Sensitivity analysis of included publications (expressed by Relative Risk and 95 % CI)

### **Publication bias**

To investigate whether the selected studies showing a more significant intervention effect than studies with null results. We plotted Funnel plot with the Trim-and-Fill adjustment and further conducted Fail-safe N method. Null study was needed to nullify the effect. We observed in the plot that the studies with higher precision tends to report higher

relative ratio. Subsequently, we performed Egger's regression test to analyze funnel plot asymmetry (Figure 5). Based on the results, we confirmed a significant publication bias ( $z = -4.3325$ ,  $p < 0.0001$ ), suggesting the size of the studies was associated with the reported risk ratio.



**Figure 5:** Funnel Plot of log Relative Risk against standard error with trim-and-fill correction

The incidence of kidney stones is increasing worldwide. Kidney stones were reported to affect approximately 1 in 11 people in the United States. The prevalence of kidney stones was 8.8 % (Scales et al., 2012). Kidney stone contributes to the development of chronic kidney disease (Rule et al., 2009) and confers threat of life of patients. The pathogenesis of kidney stone formation remains elusive and was influenced by the interplay of genetic and environmental factors (Coe et al., 1992; Curhan et al., 1997a; Moe, 2006). Risk factors such as dietary salt intake (Massey, 2005), supplemental calcium intake (Curhan et al., 1997b), diabetic mellitus (Eisner et al., 2010; Taylor et al., 2005), beverage use (Curhan et al., 1998), diet style (Taylor et al., 2009), and other lifestyle factors may affect the risk of developing kidney stone. Medical management or prevention of kidney stones has gained great attention in recent decades.

Menopause status is associated with increased urinary calcium excretion, which may increase the risk for the formation of calcium-containing kidney stones (Cappuccio et al., 2000). The occurrence of kidney stone in surgically menopausal women is higher than naturally menopausal women (Mattix Kramer et al., 2003). Hormone use and its effects on 24-hour urine composition have been reported in several laboratory and clinical studies. The results were inconsistent and sometimes controversial. For instance, in Sprague-Dawley rats treated with ethylene glycol, estrogen was found to appear to inhibit stone formation by increasing osteopontin expression in the kidneys and decreasing urinary oxalate excretion (Yagisawa et al., 2001), suggesting a potentially protective role in the incident kidney stone. In addition, estrogen use may involve in urinary excretion of kidney stone constituents and of urinary promoters. Elevated levels of serum testosterone and serum dihydrotestosterone might be involved in increased incidences of stone formation (Gupta et al., 2016). In the clinical trials conducted by Maalouf et al. (2010) PMH use is associated with higher risk of kidney stone formation

(Maalouf et al., 2010). One possible explanation for the higher incidence of stone disease with hormone therapy (HT) could be through enhanced urinary uric acid excretion with estrogen use (Adamopoulos et al., 1977; Nicholls et al., 1973). Greater uric acid excretion, in turn, could lead to heterogeneous nucleation of calcium oxalate (Pak and Arnold, 1975).

Our study has several limitations. Due to limited publications reporting the association between PMH and formation of kidney stone, only 3 studies were included in this meta-analysis. In the study of Mattix Kramer et al. (2003) lacking information of sample size in each category by different types of PMH use and menopause results in the inability of analyzing weights for each study. Different follow-up years and specifically targeted hormones may also contribute to the difficulty in identifying the true effect of PMH on the risk of kidney stone. For instance, the estrogen-only trial used conjugated equine estrogens as treatment arm while E+P trial used estrogen plus progestin. In the study of Zhao et al. (2013) they focused on measuring serum estradiol and testosterone levels in healthy postmenopausal women. Besides, we did not consider and separately analyze the incident kidney stone based on different stone components (calcium oxalate stones [COS]; non-calcium oxalate stones [NCOS]).

#### **Acknowledgements**

None.

#### **Conflict of interest**

The authors report no conflicts of interest. The authors alone are responsible for the content and writing of the paper.



## REFERENCES

- Adamopoulos D, Vlassopoulos C, Seitanides B, Contoyianis P, Vassilopoulos P. The relationship of sex steroids to uric acid levels in plasma and urine. *Acta Endocrinol.* 1977;85:198-208.
- Cappuccio FP, Kalaitzidis R, Duneclift S, Eastwood JB. Unravelling the links between calcium excretion, salt intake, hypertension, kidney stones and bone metabolism. *J Nephrol.* 2000;13:169-77.
- Coe FL, Parks JH, Asplin JR. The pathogenesis and treatment of kidney stones. *N Engl J Med.* 1992;327:1141-52.
- Curhan GC, Willett WC, Rimm EB, Stampfer MJ. Family history and risk of kidney stones. *J Am Soc Nephrol.* 1997a;8:1568-73.
- Curhan GC, Willett WC, Speizer FE, Spiegelman D, Stampfer MJ. Comparison of dietary calcium with supplemental calcium and other nutrients as factors affecting the risk for kidney stones in women. *Ann Intern Med.* 1997b;126:497-504.
- Curhan GC, Willett WC, Speizer FE, Stampfer MJ. Beverage use and risk for kidney stones in women. *Ann Intern Med.* 1998;128:534-40.
- Dey J, Creighton A, Lindberg JS, Fuselier HA, Kok DJ, Cole FE, et al. Estrogen replacement increased the citrate and calcium excretion rates in postmenopausal women with recurrent urolithiasis. *J Urol.* 2002;167:169-71.
- Eisner BH, Porten SP, Bechis SK, Stoller ML. Diabetic kidney stone formers excrete more oxalate and have lower urine pH than nondiabetic stone formers. *J Urol.* 2010;183:2244-8.
- Gupta K, Gill GS, Mahajan R. Possible role of elevated serum testosterone in pathogenesis of renal stone formation. *Int J Appl Basic Med Res.* 2016;6:241-4.
- Iguchi M, Takamura C, Umekawa T, Kurita T, Kohri K. Inhibitory effects of female sex hormones on urinary stone formation in rats. *Kidney Int.* 1999;56:479-85.
- Maalouf NM, Sato AH, Welch BJ, Howard BV, Cochrane BB, Sakhaee K, et al. Postmenopausal hormone use and the risk of nephrolithiasis: results from the Women's Health Initiative hormone therapy trials. *Arch Intern Med.* 2010;170:1678-85.
- Massey LK. Effect of dietary salt intake on circadian calcium metabolism, bone turnover, and calcium oxalate kidney stone risk in postmenopausal women. *Nutr Res.* 2005;25:891-903.
- Mattix Kramer HJ, Grodstein F, Stampfer MJ, Curhan GC. Menopause and postmenopausal hormone use and risk of incident kidney stones. *J Am Soc Nephrol.* 2003;14:1272-7.
- Moe OW. Kidney stones: pathophysiology and medical management. *Lancet.* 2006;367:333-44.
- Nicholls A, Snaith ML, Scott JT. Effect of oestrogen therapy on plasma and urinary levels of uric acid. *Br Med J.* 1973;1:449-51.
- Novak TE, Lakshmanan Y, Trock BJ, Gearhart JP, Matlaga BR. Sex prevalence of pediatric kidney stone disease in the United States: an epidemiologic investigation. *Urology.* 2009;74:104-7.
- Pak CY, Arnold LH. Heterogeneous nucleation of calcium oxalate by seeds of monosodium urate. *Proc Soc Exp Biol Med.* 1975;149:930-2.
- Rule AD, Bergstralh EJ, Melton LJ, Li X, Weaver AL, Lieske JC. Kidney stones and the risk for chronic kidney disease. *Clin J Am Soc Nephrol.* 2009;4:804-11.
- Scales CD, Smith AC, Hanley JM, Saigal CS, Project UDiA. Prevalence of kidney stones in the United States. *Eur Urol.* 2012;62:160-5.
- Taylor EN, Stampfer MJ, Curhan GC. Diabetes mellitus and the risk of nephrolithiasis. *Kidney Int.* 2005;68:1230-5.
- Taylor EN, Fung TT, Curhan GC. DASH-style diet associates with reduced risk for kidney stones. *J Am Soc Nephrol.* 2009;20:2253-9.
- Yagisawa T, Ito F, Osaka Y, Amano H, Kobayashi C, Toma H. The influence of sex hormones on renal osteopontin expression and urinary constituents in experimental urolithiasis. *J Urol.* 2001;166:1078-82.
- Zhao Z, Mai Z, Ou L, Duan X, Zeng G. Serum estradiol and testosterone levels in kidney stones disease with and without calcium oxalate components in naturally postmenopausal women. *PLoS One.* 2013;8:e75513.